Introduction to Geostatistics

9. Two-sample T-test and analysis of variance. Power and Type-II errors.

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How large should a sample be?

Given that a 95% confidence interval, e.g. for μ is obtained by

$$[\bar{X} - t_{df,\alpha}\text{SE}, \bar{X} + t_{df,\alpha}\text{SE}]$$

and given that α is chosen and σ is not under our control, we can only control the width W of the interval by manipulating n:

$$W = 2t_{df,\alpha}\text{SE} = 2t_{df,\alpha}s/\sqrt{n}$$

$$n = (\frac{2t_{df,\alpha}s}{W})^2$$

How about controlling type II errors?

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How about controlling type II errors?

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Type I and Type II errors

Of course we take a risk to wrongly rejecting a true H_0, of α.
There's however also a risk that we wrongly *not* reject a false H_0,
which is called β.

	Truth	
Test result	H_0 true	H_0 false
Reject H_0	Type I error, α	OK, $(1-\beta)$
Do not reject H_0	OK $(1-\alpha)$	Type II error, β

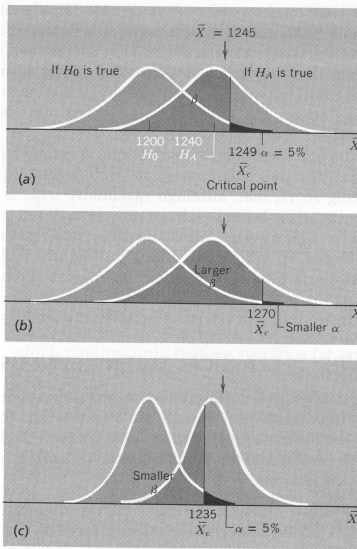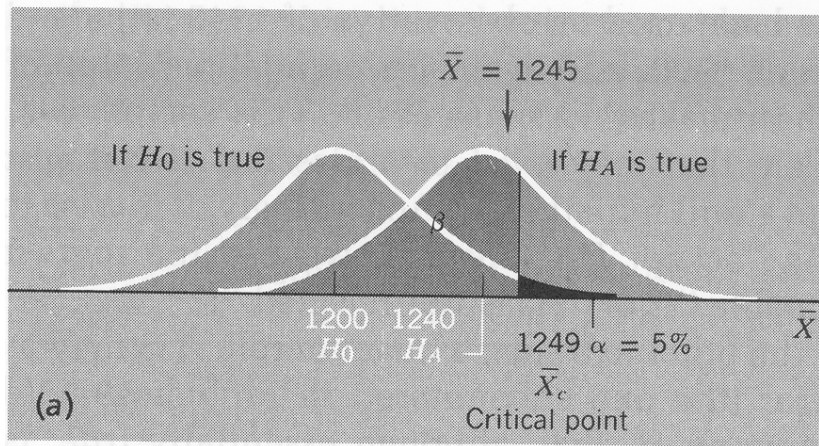Next 2 slides from: Wonnacott & Wonnacott, Introductory
statistics.

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FIGURE 9-7
(a) Hypothesis test of Figure 9-6 showing α and β. (b) How a reduction in α increases β, other things being equal. (c) How an increase in sample size allows one error probability (β) to be reduced, without increasing the other (α).

$\overline{X} = 1245$

If H_0 is true

If H_A is true

β

1200 1240
H_0 H_A

1249 $\alpha = 5\%$

$\overline{X}_c$
Critical point

$\overline{X}$

(a)

How to compute the power function?

- Given that H_0 is not true, then what is true? Probabilities cannot be computed without assumptions about the population.
- Given a fixed H_A, we can compute power as in the figure in the previous slide.
- For all possible H_A's, we obtain the *power function*.
- What determines the power?

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- What determines the power?
  - The difference between the H_0 and H_A means (delta)
  - The width of the curves ($SE = \sigma/\sqrt{n}$)
  - α
  - where is α? – one-sided or two-sided
  - what is n? how is SE computed? – type of test: one-sample, two-sample, paired

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Power computation using `power.t.test`

Description: Compute power of test, or determine parameters to obtain target power.

Details: Exactly one of the parameters n, `delta`, `power`, `sd`, and `sig.level` must be passed as NULL, and that parameter is determined from the others. Notice that the last two have non-NULL defaults so NULL must be explicitly passed if you want to compute them.

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Compute sample size

```
> power.t.test(n = NULL, delta = 1, sd = 1, sig.level = 0.05,
+     power = 0.9, type = "two.sample", alternative = "two.sided")

     Two-sample t test power calculation

              n = 22.02110
          delta = 1
             sd = 1
      sig.level = 0.05
          power = 0.9
    alternative = two.sided

 NOTE: n is number in *each* group
```

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Compute delta (H_A)

```
> power.t.test(n = 20, delta = NULL, sd = 1, sig.level = 0.05,
+     power = 0.9, type = "two.sample", alternative = "two.sided")

     Two-sample t test power calculation

              n = 20
          delta = 1.051970
             sd = 1
      sig.level = 0.05
          power = 0.9
    alternative = two.sided

 NOTE: n is number in *each* group
```

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Compute power

```
> power.t.test(n = 20, delta = 1, sd = 1, sig.level = 0.05,
+     power = NULL, type = "two.sample", alternative = "two.sided")

     Two-sample t test power calculation

              n = 20
          delta = 1
             sd = 1
      sig.level = 0.05
          power = 0.8689528
    alternative = two.sided

 NOTE: n is number in *each* group
```

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Compute significance level

```
> power.t.test(n = 20, delta = 1, sd = 1, sig.level = NULL,
+     power = 0.9, type = "two.sample", alternative = "two.sided")

     Two-sample t test power calculation

              n = 20
          delta = 1
             sd = 1
      sig.level = 0.07004584
          power = 0.9
    alternative = two.sided

 NOTE: n is number in *each* group
```
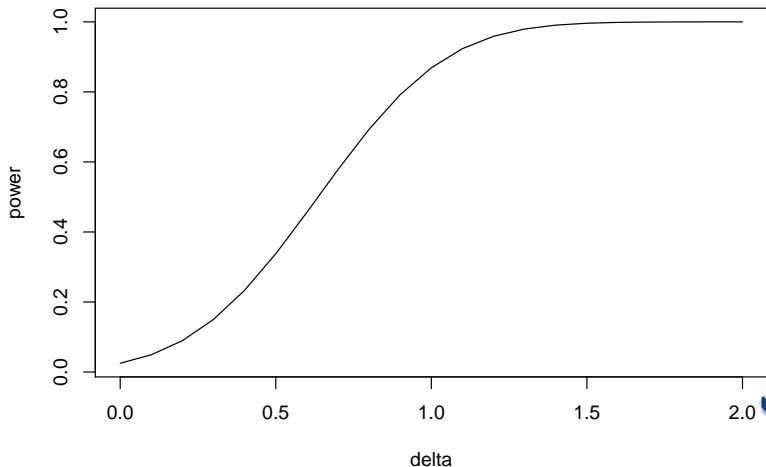
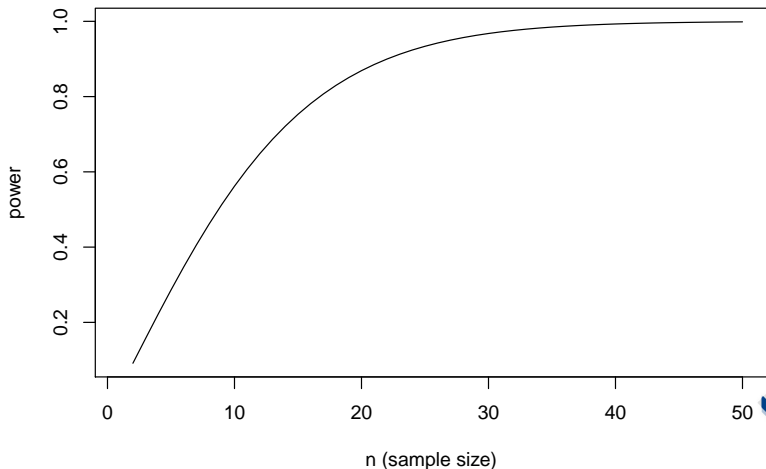(Note that this is of little operational use; computing sd is of even less operational use)

Compute power function vs. delta, n = 20, s = 1

```
> plot((0:20)/10, power.t.test(power = NULL, delta = (0:20)/10,
+     n = 20)$power, type = "l", xlab = "delta", ylab = "power")
```

Compute power function vs. n; delta = 1

```
> plot(1:50, power.t.test(delta = 1, n = 1:50)$power, type = "l",
+      xlab = "n (sample size)", ylab = "power")
```

The power concept beyond n

In a testing framework, increasing n will make every small difference in means significant, as small differences will be noted (with large power). This does not mean that the difference found is relevant.

Suppose we're studying the effect of a medication type on health, or a herbicide type on plant disease. Two large samples (with and without treatment) confirmed (showed significantly) that in the group without treatment there was 45% succes, less than in the group with treatment with 47% success.

That's OK, but should we now collectively apply the treatment? Do the effects compensate for the costs and side effects?

Significance is something else as relevance

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Significance is something else as relevance

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The power concept beyond n

Taking a larger sample always increases power. Can we do something else to increase power? Yes: choose a more appropriate analysis. Recall the paired data of lecture 7:

obj	t_1	t_2
1	13.5	12.7
2	15.3	15.1
3	7.5	6.6
4	10.3	8.5
5	8.7	8.0

```
> x1 = c(13.5, 15.3, 7.5, 10.3, 8.7)
> x2 = c(12.7, 15.1, 6.6, 8.5, 8)
> x1 - x2

[1] 0.8 0.2 0.9 1.8 0.7
```

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```
> t.test(x1, x2, var.equal = TRUE)

        Two Sample t-test

data:  x1 and x2
t = 0.4066, df = 8, p-value = 0.695
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -4.111314  5.871314
sample estimates:
mean of x mean of y
    11.06     10.18

> t.test(x1, x2, paired = TRUE)

        Paired t-test

data:  x1 and x2
t = 3.3896, df = 4, p-value = 0.02754
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 0.1591929 1.6008071
sample estimates:
mean of the differences
                   0.88
```

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```
> power.t.test(delta = 0.88, n = 5, sd = sqrt((var(x1) +
+     var(x2))/2))

     Two-sample t test power calculation

              n = 5
          delta = 0.88
             sd = 3.422353
      sig.level = 0.05
          power = 0.0548756
    alternative = two.sided

 NOTE: n is number in *each* group

> power.t.test(delta = 0.88, n = 5, sd = sd(x1 - x2), type = "paired")

     Paired t test power calculation

              n = 5
          delta = 0.88
             sd = 0.580517
      sig.level = 0.05
          power = 0.7192318
    alternative = two.sided
```
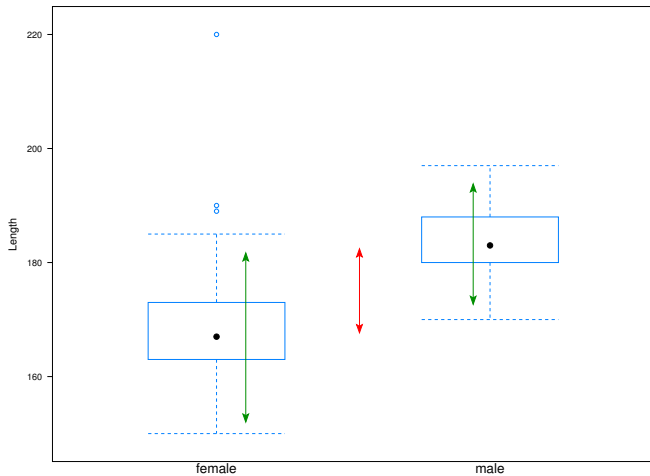


```
 NOTE: n is number of *pairs*, sd is std.dev. of *differences* within pa
```

Two-sample T-test and analysis of variance.

$$t = \frac{\bar{X}_1 - \bar{X}_2}{s/\sqrt{n}}$$

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Generalizing 2 to p groups: from t to F

- let $n_1 = n_2 = n$ and $N = n_1 + n_2$, and assume $\sigma_1 = \sigma_2$
- $t = \frac{\bar{X}_1 - \bar{X}_2}{s/\sqrt{n}} = \frac{(\bar{X}_1 - \bar{X}_2)\sqrt{n}}{s}$
- $t^2 = \frac{(\bar{X}_1 - \bar{X}_2)^2 n}{s^2} = \frac{n\text{Var}(\bar{X}_i)}{s^2}$
- $\frac{n\text{Var}(\bar{X}_i)}{s^2}$, with s^2 the pooled (averaged, joined) within-group variance
- numerator: variance, as obtained from variability between groups (group means)
- denominator: variance, as obtained from variability within groups (ignores differences between groups)
- Under the hypothesis $H_0 : \mu_1 = \mu_2 = ... = \mu_p$,

$$F = \frac{n\text{Var}(\bar{X}_i)}{s^2}$$

follows the F distribution with $p - 1$ (numerator) and $N - p$ (denominator) degrees of freedom.

This idea generalizes the two-sample t-test, testing $H_0 : \mu_1 = \mu_2$ to the F-test, testing $H_0 : \mu_1 = \mu_2 = ... = \mu_p$.

Generalizing 2 to p groups: from t to F

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Why not use many t-tests?

- Suppose we have three groups, and we can reject
  $H_0 : \mu_1 = \mu_2$, we can reject $H_0 : \mu_2 = \mu_3$, but cannot reject
  $H_0 : \mu 1 = \mu_3$.
  This will be clumsy to explain.
  When hypothesis $H_0 : \mu_1 = \mu_2 = \mu_3$ is rejected, we can
  continue work under the hypothesis "the group means are not
  identical".
- Suppose we have many (10) groups with few observations (3)
  each. Pairwise testing has very little power (df = 4), whereas
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How to read ANOVA tables?

```
> summary(aov(Length ~ Gender))
            Df Sum Sq Mean Sq F value    Pr(>F)
Gender       1 4017.1  4017.1  45.466 1.845e-09 ***
Residuals   84 7421.8    88.4
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Gender: effect, explanatory variable, grouping variable, between groups

Residuals: error, within-groups, unexplained variability
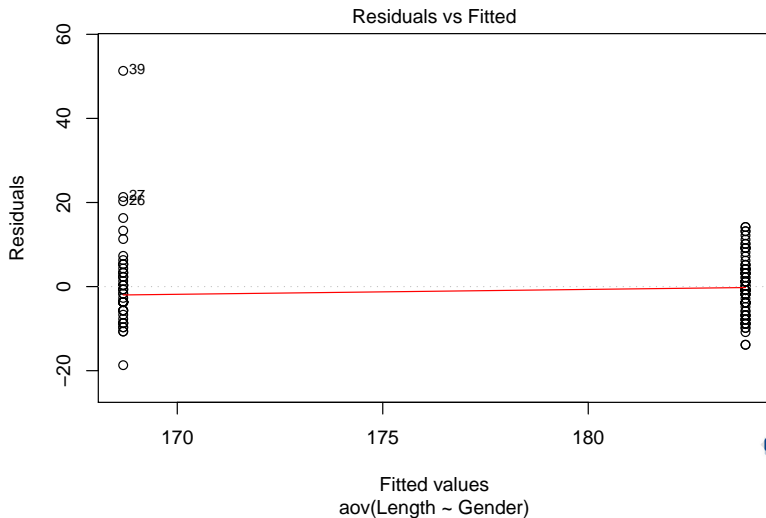
Df: degrees of freedom for that row

Sum Sq: sum of squares, between or within
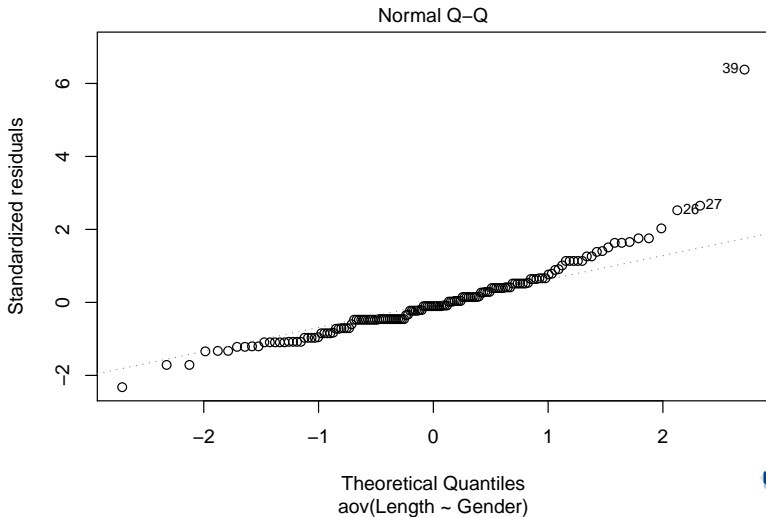
Mean Sq: mean squares: Sum Sq divided by Df

F value: Mean Sq effect divided by Mean Sq Residuals
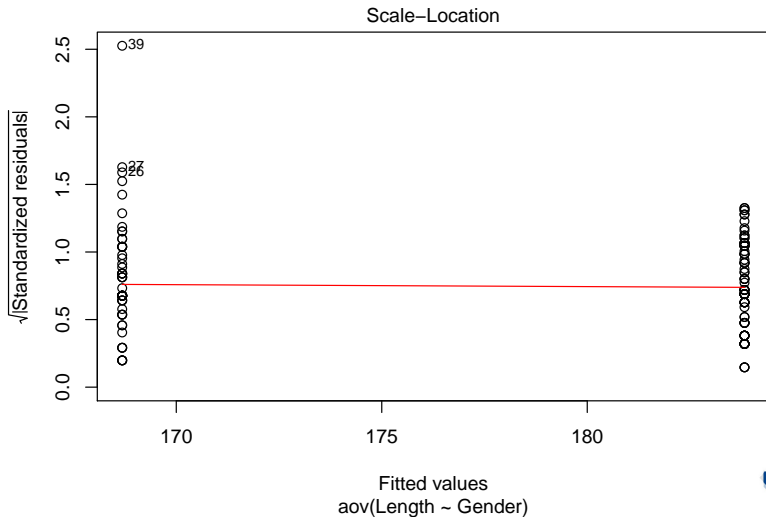
Pr(>F): significance level, p-value

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```
> plot(aov(Length ~ Gender), which = 1)
> Length[31]
```

`[1] 162`

Residuals vs Fitted

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```
> plot(aov(Length ~ Gender), which = 2)
```


Normal Q–Q

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```
> plot(aov(Length ~ Gender), which = 3)
```


Scale–Location

Two-way ANOVA; setting up data

The data can also be organized like this:

```
> x = data.frame(resp = c(x1, x2), time = rep(c("t1", "t2"),
+     each = 5), obj = rep(letters[1:5], 2))
> x

   resp time obj
1  13.5   t1   a
2  15.3   t1   b
3   7.5   t1   c
4  10.3   t1   d
5   8.7   t1   e
6  12.7   t2   a
7  15.1   t2   b
8   6.6   t2   c
9   8.5   t2   d
10  8.0   t2   e
```

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Two-way ANOVA

One-way ANOVA:

```
> summary(aov(resp ~ time, x))

            Df Sum Sq Mean Sq F value Pr(>F)
time         1  1.936   1.936  0.1653  0.695
Residuals    8 93.700  11.713
```

Two-way ANOVA:

```
> summary(aov(resp ~ time + obj, x))

            Df Sum Sq Mean Sq F value    Pr(>F)
time         1  1.936  1.9360  11.490 0.0275393 *
obj          4 93.026 23.2565 138.021 0.0001545 ***
Residuals    4  0.674  0.1685
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

ifgi

Paired t-test vs two-way ANOVA

Paired t-test:

```
> t.test(x1, x2, paired = TRUE)

        Paired t-test

data:  x1 and x2
t = 3.3896, df = 4, p-value = 0.02754
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 0.1591929 1.6008071
sample estimates:
mean of the differences
                   0.88
```

Two-way ANOVA:

```
> summary(aov(resp ~ time + obj, x))

            Df Sum Sq Mean Sq F value    Pr(>F)
time         1  1.936  1.9360  11.490 0.0275393 *
obj          4 93.026 23.2565 138.021 0.0001545 ***
Residuals    4  0.674  0.1685
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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Paired t-test vs. two-way ANOVA

Note: p-values are identical. Anova generalizes paired t-tests in the sense that e.g. time can have more than 2 levels (but is considered categorical).
Further extensions: three-way, more-way anova; interactions.
Now introduce the `meuse` data set

ifgi

Paired t-test vs. two-way ANOVA

Note: *p*-values are identical. Anova generalizes paired t-tests in the sense that e.g. time can have more than 2 levels (but is considered categorical).

Further extensions: three-way, more-way anova; interactions.

Now introduce the `meuse` data set

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